

On Optimal Hypothesis Testing for Pair of Stochastically Coupled Objects.

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Abstract

The paper is devoted to hypotheses testing for a model consisting of two stochastically coupled objects. It is supposed that L_1 possible probability distributions are known for the first object and the second object is distributed according to one of $L_1 \times L_2$ given conditional distributions depending on the distribution index and the current observed state of the first object. The matrix of interdependencies of all possible pairs of the error probability exponents in asymptotically optimal tests of distributions of both objects is studied. The case of two objects which cannot have the same probability distribution from two possible variants was considered by Ahlswede and Haroutunian. This case for three hypotheses and the model of two statistically dependent objects for two hypotheses were examined by Haroutunian and Yessayan.

1 Problem Statement and Preliminary Results

In this paper we solve a generalisation of the problem of many hypotheses testing concerning one object [1]. In paper [2] Ahlswede and Haroutunian and in [3] Haroutunian formulated a number of problems on multiple hypotheses testing and identification. Haroutunian and Hakobyan in [4] examined the models of two independent objects with three and in [5] with L hypotheses. Haroutunian and Yessayan in [6] studied the model of two objects which can have only different distributions from possible three.

Let X_1 and X_2 be random variables (RVs) taking values in a finite set \mathcal{X} and $\mathcal{P}(\mathcal{X})$ be the space of all possible distributions on \mathcal{X} . There are given L_1 probability distributions (PD) $G_{l_1} = \{G_{l_1}(x^1), x^1 \in \mathcal{X}\}$, $l_1 = \overline{1, L_1}$, from $\mathcal{P}(\mathcal{X})$. The first object characterized by RV X_1 can have one of these L_1 distributions and the second object dependent on the first, and characterized by RV X_2 can have one of $L_1 \times L_2$ conditional PDs $G_{l_2/l_1} = \{G_{l_2/l_1}(x^2|x^1), x^1, x^2 \in \mathcal{X}\}$, $l_1 = \overline{1, L_1}$, $l_2 = \overline{1, L_2}$. Let $(x_1, x_2) = ((x_1^1, x_1^2), (x_2^1, x_2^2), \dots, (x_N^1, x_N^2))$ be a sequence of results of N independent observations of the vector (X_1, X_2) , which can have one of $L_1 \times L_2$ joint PDs $G_{l_1, l_2}(x^1, x^2)$, $l_1 = \overline{1, L_1}$, $l_2 = \overline{1, L_2}$, where $G_{l_1, l_2}(x^1, x^2) = G_{l_1}(x^1)G_{l_2/l_1}(x^2|x^1)$. The probability of vector (x_1, x_2) is defined by PD G_{l_1, l_2}

$$G_{l_1, l_2}^N(x_1, x_2) = G_{l_1}^N(x_1)G_{l_2/l_1}^N(x_2|x_1) = \prod_{n=1}^N G_{l_1}(x_n^1)G_{l_2/l_1}(x_n^2|x_n^1).$$

where $G_{l_1}^N(\mathbf{x}_1) = \prod_{n=1}^N G_{l_1}(x_1^n)$ and $G_{l_2/l_1}^N(\mathbf{x}_2|\mathbf{x}_1) = \prod_{n=1}^N G_{l_2/l_1}(x_2^n|x_1^n)$.

The test, which we denote by Φ^N , is a procedure of making decision about indices of distributions on the base of N observations of objects. The test Φ^N may be composed of a pair of tests φ_1^N and φ_2^N for the separate objects: $\Phi^N = (\varphi_1^N, \varphi_2^N)$. For the object characterized by X_1 the non-randomized test $\varphi_1^N(\mathbf{x}_1)$ can be determined by partition of the sample space \mathcal{X}^N on L_1 disjoint subsets $\mathcal{A}_{l_1}^N = \{\mathbf{x}_1 : \varphi_1^N(\mathbf{x}_1) = l_1\}$, $l_1 = \overline{1, L_1}$, i.e. the set $\mathcal{A}_{l_1}^N$ consists of vectors \mathbf{x}_1 for which the PD G_{l_1} is adopted. The probability $\alpha_{l_1|m_1}^N(\varphi_1^N)$ of the erroneous acceptance of PD G_{l_1} , provided that G_{m_1} is true, $l_1, m_1 = \overline{1, L_1}$, $m_1 \neq l_1$, is defined by the set $\mathcal{A}_{l_1}^N$

$$\alpha_{l_1|m_1}^N(\varphi_1^N) \triangleq G_{m_1}^N(\mathcal{A}_{l_1}^N). \quad (1)$$

We define the probability to reject G_{m_1} , when it is true, as follows

$$\alpha_{m_1|m_1}^N(\varphi_1^N) \triangleq \sum_{l_1 \neq m_1} \alpha_{l_1|m_1}^N(\varphi_1^N) = G_{m_1}^N(\overline{\mathcal{A}_{m_1}^N}). \quad (2)$$

Denote by φ_1, φ_2 and Φ the infinite sequences of tests. Corresponding error probability exponents $E_{l_1|m_1}(\varphi_1)$ for test φ_1 called reliabilities are defined as

$$E_{l_1|m_1}(\varphi_1) \triangleq \overline{\lim}_{N \rightarrow \infty} -\frac{1}{N} \log \alpha_{l_1|m_1}^N(\varphi_1^N), \quad m_1, l_1 = \overline{1, L_1}. \quad (3)$$

It follows from (2) and (3) that

$$E_{m_1|m_1}(\varphi_1) = \min_{l_1 \neq m_1} E_{l_1|m_1}(\varphi_1), \quad l_1, m_1 = \overline{1, L_1}, \quad l_1 \neq m_1. \quad (4)$$

For the second object characterized by RV X_2 the non-randomized test $\varphi_2^N(\mathbf{x}_2, \mathbf{x}_1, l_1)$ depending on vectors $\mathbf{x}_1, \mathbf{x}_2$ and on the index of the hypothesis l_1 adopted for X_1 , can be given by division of the sample space \mathcal{X}^N on L_2 disjoint subsets $\mathcal{A}_{l_2/l_1}^N(\mathbf{x}_1) = \{\mathbf{x}_2 : \varphi_2^N(\mathbf{x}_2, \mathbf{x}_1, l_1) = l_2\}$, $l_1 = \overline{1, L_1}, l_2 = \overline{1, L_2}$. The set $\mathcal{A}_{l_2/l_1}^N(\mathbf{x}_1)$ consists of vectors \mathbf{x}_2 for which the PD G_{l_2/l_1} is adopted. The probabilities of the erroneous acceptance of PD G_{l_2/l_1} , provided that G_{m_2/m_1} is true are the following

$$\alpha_{l_2/l_1, m_1, m_2}^N(\varphi_2^N) \triangleq G_{m_2/l_1}^N(\mathcal{A}_{l_2/l_1}^N(\mathbf{x}_1)|\mathbf{x}_1), \quad l_1, m_1 = \overline{1, L_1}, \quad l_2, m_2 = \overline{1, L_2}, \quad m_2 \neq l_2. \quad (5)$$

The corresponding reliabilities, are defined as

$$E_{l_2/l_1, m_1, m_2}(\varphi_2) \triangleq \overline{\lim}_{N \rightarrow \infty} -\frac{1}{N} \log \alpha_{l_2/l_1, m_1, m_2}^N(\varphi_2^N), \quad l_1, m_1 = \overline{1, L_1}, \quad l_2, m_2 = \overline{1, L_2}, \quad m_2 \neq l_2. \quad (6)$$

It is clear from (5) and (6) that

$$E_{m_2/l_1, m_1, m_2}(\varphi_2) = \min_{l_2 \neq m_2} E_{l_2/l_1, m_1, m_2}(\varphi_2), \quad l_1, m_1 = \overline{1, L_1}, \quad l_2, m_2 = \overline{1, L_2}. \quad (7)$$

The matrices $\mathbf{E}(\varphi_1) = \{E_{l_1|m_1}(\varphi_1), l_1, m_1 = \overline{1, L_1}\}$, $\mathbf{E}(\varphi_2) = \{E_{l_2/l_1, m_1, m_2}(\varphi_2), l_1, m_1 = \overline{1, L_1}, l_2, m_2 = \overline{1, L_2}\}$ are called the reliability matrices of the sequence of tests φ_1, φ_2 . For two objects we study the probabilities $\alpha_{l_1, l_2|m_1, m_2}^N(\Phi^N)$ of the erroneous acceptance by the test Φ^N of the pair of PDs $G_{l_1}, G_{l_2/l_1}$ (or joint PD G_{l_1, l_2}) provided that the pair

$G_{m_1}, G_{m_2/l_1}$ (or joint PD G_{m_1, m_2}) is true, where $(m_1, m_2) \neq (l_1, l_2)$, $l_1, m_1 = \overline{1, L_1}$, $l_2, m_2 = \overline{1, L_2}$. The probability to reject a true PD G_{m_1, m_2} , is defined as follows

$$\alpha_{m_1, m_2 | m_1, m_2}^N(\Phi^N) \triangleq \sum_{(l_1, l_2) \neq (m_1, m_2)} \alpha_{l_1, l_2 | m_1, m_2}^N(\Phi^N), \quad l_1, m_1 = \overline{1, L_1}, \quad l_2, m_2 = \overline{1, L_2}. \quad (8)$$

The reliabilities of the sequence of tests Φ are the following

$$E_{l_1, l_2 | m_1, m_2}(\Phi) \triangleq \lim_{N \rightarrow \infty} - \frac{1}{N} \log \alpha_{l_1, l_2 | m_1, m_2}^N(\Phi^N), \quad l_1, m_1 = \overline{1, L_1}, \quad l_2, m_2 = \overline{1, L_2}. \quad (9)$$

From (8) and (9) we have

$$E_{m_1, m_2 | m_1, m_2}(\Phi) = \min_{(l_1, l_2) \neq (m_1, m_2)} E_{l_1, l_2 | m_1, m_2}(\Phi), \quad l_1, m_1 = \overline{1, L_1}, \quad l_2, m_2 = \overline{1, L_2}. \quad (10)$$

We call the matrix $E(\Phi) = \{E_{l_1, l_2 | m_1, m_2}(\Phi), l_1, m_1 = \overline{1, L_1}, l_2, m_2 = \overline{1, L_2}\}$ the reliability matrix of the sequence of tests Φ . Our aim is to study the reliability matrix of optimal tests, and the conditions of positivity of all its elements.

Definition: We call the sequence of tests φ_1^* (or φ_2^* , or Φ^*) logarithmically asymptotically optimal (LAO) if for given positive values of $L_1 - 1$ (or $L_2 - 1$, or $(L_1 - 1)(L_2 - 1)$) diagonal elements of the corresponding matrix $E(\varphi_1^*)$ (or $E(\varphi_2^*)$, or $E(\Phi^*)$) maximal values to all other elements of it are provided.

The following lemma is an extension of the lemmas from [2] and [4].

Lemma: If the reliabilities $E_{l_1 | m_1}$ and $E_{l_2 | l_1, m_1, m_2}$ of tests φ_1^N and φ_2^N are strictly positive, then the following relations hold:

$$E_{l_1, l_2 | m_1, m_2} = E_{l_1 | m_1} + E_{l_2 | l_1, m_1, m_2}, \quad \text{for } m_1 \neq l_1, m_2 \neq l_2, \quad (11.a)$$

$$E_{l_1, l_2 | m_1, m_2} = E_{l_1 | m_1}, \quad \text{for } m_1 \neq l_1, m_2 = l_2, \quad (11.b)$$

$$E_{l_1, l_2 | m_1, m_2} = E_{l_2 | l_1, m_1, m_2}, \quad \text{for } m_1 = l_1, m_2 \neq l_2. \quad (11.c)$$

Proof: The following equalities are valid for error probabilities:

$$\alpha_{l_1, l_2 | m_1, m_2}^N = \alpha_{l_1 | m_1}^N \alpha_{l_2 | l_1, m_1, m_2}^N, \quad \text{for } m_1 \neq l_1, m_2 \neq l_2, \quad (12.a)$$

$$\alpha_{l_1, l_2 | m_1, m_2}^N = \alpha_{l_1 | m_1}^N (1 - \alpha_{l_2 | l_1, m_1, m_2}^N), \quad \text{for } m_1 \neq l_1, m_2 = l_2, \quad (12.b)$$

$$\alpha_{l_1, l_2 | m_1, m_2}^N = (1 - \alpha_{l_1 | m_1}^N) \alpha_{l_2 | l_1, m_1, m_2}^N, \quad \text{for } m_1 = l_1, m_2 \neq l_2. \quad (12.c)$$

Thus, in light of (3), (6) and (9), we can obtain (11).

We shall reformulate now the Theorem from [1] for the case of one object with L_1 hypotheses. This requires some notions and notations. We define the entropy $H_{Q_{x_1}}(X_1)$ and the informational divergence $D(Q_{x_1} \| G_{l_1})$, $l_1 = \overline{1, L_1}$, as follows:

$$H_{Q_{x_1}}(X_1) \triangleq - \sum_{x^1 \in \mathcal{X}} Q_{x_1}(x^1) \log Q_{x_1}(x^1),$$

$$D(Q_{x_1} \| G_{l_1}) \triangleq \sum_{x^1 \in \mathcal{X}} Q_{x_1}(x^1) \log \frac{Q_{x_1}(x^1)}{G_{l_1}(x^1)}.$$

For given positive numbers $E_{1|1}, \dots, E_{L-1|L-1}$, let us consider the following sets of PDs $Q = \{Q(x^1), x^1 \in \mathcal{X}\}$:

$$\mathcal{R}_{l_1} \triangleq \{Q: D(Q||G_{l_1}) \leq E_{l_1|l_1}\}, \quad l_1 = \overline{1, L_1 - 1}, \quad (13a)$$

$$\mathcal{R}_{L_1} \triangleq \{Q: D(Q||G_{L_1}) > E_{L_1|L_1}\}, \quad l_1 = \overline{1, L_1 - 1}, \quad (13b)$$

and the elements of the reliability matrix \mathbf{E}^* of the LAO test:

$$E_{l_1|l_1}^* = E_{l_1|l_1}^*(E_{l_1|l_1}) \triangleq E_{l_1|l_1}, \quad l_1 = \overline{1, L_1 - 1}, \quad (14a)$$

$$E_{l_1|m_1}^* = E_{l_1|m_1}^*(E_{l_1|l_1}) \triangleq \inf_{Q \in \mathcal{R}_{m_1}} D(Q||G_{m_1}), \quad m_1 = \overline{1, L_1}, \quad m_1 \neq l_1, \quad l_1 = \overline{1, L_1 - 1}, \quad (14b)$$

$$E_{L_1|m_1}^* = E_{L_1|m_1}^*(E_{1|1}, E_{2|2}, \dots, E_{L_1-1|L_1-1}) \triangleq \inf_{Q \in \mathcal{R}_{m_1}} D(Q||G_{m_1}), \quad m_1 = \overline{1, L_1 - 1}, \quad (14c)$$

$$E_{L_1|L_1}^* = E_{L_1|L_1}^*(E_{1|1}, E_{2|2}, \dots, E_{L_1-1|L_1-1}) \triangleq \min_{l_1 = \overline{1, L_1 - 1}} E_{l_1|L_1}^*, \quad (14d)$$

Theorem 1[1]: If all distributions G_{l_1} , $l_1 = \overline{1, L_1}$, are different in the sense that $D(G_{l_1}||G_{m_1}) > 0$, $l_1 \neq m_1$, and the positive numbers $E_{1|1}, E_{2|2}, \dots, E_{L_1-1|L_1-1}$ are such that the following inequalities hold

$$E_{1|1} < \min_{l_1 = \overline{2, L_1}} D(G_{l_1}||G_1), \quad (15)$$

$$E_{m_1|m_1} < \min_{l_1 = m_1 + 1, L_1} D(G_{l_1}||G_{m_1}), \quad \min_{l_1 = 1, m_1 - 1} E_{l_1|m_1}^*(E_{l_1|l_1}), \quad m_1 = \overline{2, L_1 - 1},$$

then there exists a LAO sequence of tests φ_i^* , the reliability matrix of which $\mathbf{E}^* = \{E_{m_1|l_1}(\varphi_i^*)\}$ is defined in (14) and all elements of it are positive.

When one of the inequalities (15) is violated, then at least one element of the matrix \mathbf{E}^* is equal to 0.

2 LAO Testing of L_2 Hypotheses for the Second (Dependent) Object

We need some notions and estimates from the method of types [7] - [10]. The type of a vector \mathbf{x}_1 is a PD

$$Q_{\mathbf{x}_1} = \{Q_{\mathbf{x}_1}(x^1) = \frac{1}{N} N(x^1|\mathbf{x}_1), x^1 \in \mathcal{X}\},$$

where $N(x^1|\mathbf{x}_1)$ is the number of repetitions of the symbol x^1 in vector \mathbf{x}_1 . The subset of $\mathcal{P}(\mathcal{X})$ consisting of the possible types of sequences $\mathbf{x}_1 \in \mathcal{X}$ is denoted by $\mathcal{P}_N(\mathcal{X})$. The set of all vectors \mathbf{x}_1 of the type $Q_{\mathbf{x}_1}$ is denoted by $T_{Q_{\mathbf{x}_1}}^N(X_1)$, remark that $T_{Q_{\mathbf{x}_1}}^N(X_1) = \emptyset$ for $Q \notin \mathcal{P}_N(\mathcal{X})$. The following estimates for the set $T_{Q_{\mathbf{x}_1}}^N(X_1)$ of all vectors of the same type $Q_{\mathbf{x}_1}$ hold

$$(N+1)^{-|\mathcal{X}|} \exp\{NH_{Q_{\mathbf{x}_1}}(X_1)\} \leq |T_{Q_{\mathbf{x}_1}}^N(X_1)| \leq \exp\{NH_{Q_{\mathbf{x}_1}}(X_1)\}.$$

For a pair of sequences $(\mathbf{x}_1, \mathbf{x}_2) \in \mathcal{X}^N \times \mathcal{X}^N$ let $N(x^1, x^2|\mathbf{x}_1, \mathbf{x}_2)$ be the number of occurrences of pair $(x^1, x^2) \in \mathcal{X} \times \mathcal{X}$ in the same places of the pair of vectors $(\mathbf{x}_1, \mathbf{x}_2)$. The joint type of the pair $(\mathbf{x}_1, \mathbf{x}_2)$ is PD $Q_{\mathbf{x}_1, \mathbf{x}_2} = \{Q_{\mathbf{x}_1, \mathbf{x}_2}(x^1, x^2), x^1, x^2 \in \mathcal{X}\}$ defined by

$$Q_{\mathbf{x}_1, \mathbf{x}_2}(x^1, x^2) \triangleq \frac{1}{N} N(x^1, x^2|\mathbf{x}_1, \mathbf{x}_2), \quad x^1, x^2 \in \mathcal{X}.$$

The conditional type of \mathbf{x}_2 for given \mathbf{x}_1 is the conditional distribution $V_{\mathbf{x}_1, \mathbf{x}_2} \triangleq \{V_{\mathbf{x}_1, \mathbf{x}_2}(x^2|x^1), x^1, x^2 \in \mathcal{X}\}$ defined by

$$V_{\mathbf{x}_1, \mathbf{x}_2}(x^2|x^1) \triangleq \frac{Q_{\mathbf{x}_1, \mathbf{x}_2}(x^1, x^2)}{Q_{\mathbf{x}_1}(x^1)} = \frac{N(x^1, x^2|\mathbf{x}_1, \mathbf{x}_2)}{N(x^1|\mathbf{x}_1)}, \quad x^1, x^2 \in \mathcal{X}.$$

The conditional entropy of RV X_2 for given X_1 is:

$$H_{Q_{\mathbf{x}_1}, V_{\mathbf{x}_1, \mathbf{x}_2}}(X_2 | X_1) = - \sum_{x^1, x^2} Q_{\mathbf{x}_1}(x^1) V_{\mathbf{x}_1, \mathbf{x}_2}(x^2|x^1) \log V_{\mathbf{x}_1, \mathbf{x}_2}(x^2|x^1).$$

For some conditional PD $V = \{V(x^2|x^1), x^1, x^2 \in \mathcal{X}\}$ the conditional divergences of PD $\{Q_{\mathbf{x}_1}(x^1) V(x^2|x^1), x^1, x^2 \in \mathcal{X}\}$ with respect to PD $\{Q_{\mathbf{x}_1}(x^1) G_{l_2/l_1}(x^2|x^1), x^1, x^2 \in \mathcal{X}\}$ for all l_1, l_2 is defined as follows

$$D(V||G_{l_2/l_1}|Q_{\mathbf{x}_1}) \triangleq \sum_{x^1, x^2} Q_{\mathbf{x}_1}(x^1) V(x^2|x^1) \log \frac{V(x^2|x^1)}{G_{l_2/l_1}(x^2|x^1)},$$

also

$$D(G_{l_2/l_1}||G_{m_2/l_1}|Q_{\mathbf{x}_1}) \triangleq \sum_{x^1, x^2} Q_{\mathbf{x}_1}(x^1) G_{l_2/l_1}(x^2|x^1) \log \frac{G_{l_2/l_1}(x^2|x^1)}{G_{m_2/l_1}(x^2|x^1)}.$$

The family of vectors \mathbf{x}_2 of the conditional type $V_{\mathbf{x}_1, \mathbf{x}_2}$ for given \mathbf{x}_1 of the type $Q_{\mathbf{x}_1}$ is denoted by $T_{Q_{\mathbf{x}_1}, V_{\mathbf{x}_1, \mathbf{x}_2}}^N(X_2 | \mathbf{x}_1)$ and called V -shell of \mathbf{x}_1 . The set of all possible V -shells for \mathbf{x}_1 of type $Q_{\mathbf{x}_1}$ is denoted by $\mathcal{V}_N(\mathcal{X}, Q_{\mathbf{x}_1})$. For any conditional type $V_{\mathbf{x}_1, \mathbf{x}_2}$ and $\mathbf{x}_1 \in T_{Q_{\mathbf{x}_1}}^N(X_1)$ it is known that

$$(N+1)^{-|\mathcal{X}|^2} \exp\{NH_{Q_{\mathbf{x}_1}, V_{\mathbf{x}_1, \mathbf{x}_2}}(X_2|X_1)\} \leq |T_{Q_{\mathbf{x}_1}, V_{\mathbf{x}_1, \mathbf{x}_2}}^N(X_2 | \mathbf{x}_1)| \leq \exp\{NH_{Q_{\mathbf{x}_1}, V_{\mathbf{x}_1, \mathbf{x}_2}}(X_2|X_1)\}. \quad (16)$$

For given positive numbers $E_{l_2|l_1, m_1, l_2}$, $l_2 = \overline{1, L_2 - 1}$, for each pair $l_1, m_1 = \overline{1, L_1}$ and for $Q_{\mathbf{x}_1}$ let us define the following regions and values:

$$\mathcal{R}_{l_2/l_1}(Q_{\mathbf{x}_1}) \triangleq \{V : D(V||G_{l_2/l_1}|Q_{\mathbf{x}_1}) \leq E_{l_2|l_1, m_1, l_2}\}, \quad l_2 = \overline{1, L_2 - 1}, \quad (17.a)$$

$$\mathcal{R}_{L_2/l_1}(Q_{\mathbf{x}_1}) \triangleq \{V : D(V||G_{L_2/l_1}|Q_{\mathbf{x}_1}) > E_{L_2|l_1, m_1, L_2}\}, \quad l_2 = \overline{1, L_2 - 1}, \quad (17.b)$$

$$E_{l_2|l_1, m_1, l_2}^* = E_{l_2|l_1, m_1, l_2}^*(E_{l_2|l_1, m_1, l_2}) \triangleq E_{l_2|l_1, m_1, l_2}, \quad l_2 = \overline{1, L_2 - 1}, \quad (18.a)$$

$$E_{l_2|l_1, m_1, m_2}^* = E_{l_2|l_1, m_1, m_2}^*(E_{l_2|l_1, m_1, l_2}) \triangleq \inf_{V \in \mathcal{R}_{l_2/l_1}(Q_{\mathbf{x}_1})} D(V||G_{m_2/l_1}|Q_{\mathbf{x}_1}),$$

$$m_2 = \overline{1, L_2}, \quad m_2 \neq l_2, \quad l_2 = \overline{1, L_2 - 1}, \quad (18.b)$$

$$E_{L_2|l_1, m_1, m_2}^* = E_{L_2|l_1, m_1, m_2}^*(E_{l_1|l_1, m_1, 1}, E_{l_2|l_1, m_1, 2}, \dots, E_{L_2-1|l_1, m_1, L_2-1}) \triangleq$$

$$\triangleq \inf_{V \in \mathcal{R}_{L_2/l_1}(Q_{\mathbf{x}_1})} D(V||G_{m_2/l_1}|Q_{\mathbf{x}_1}), \quad m_2 = \overline{1, L_2 - 1}, \quad (18.c)$$

$$E_{L_2|l_1, m_1, L_2}^* = E_{L_2|l_1, m_1, L_2}^*(E_{l_1|l_1, m_1, 1}, E_{l_2|l_1, m_1, 2}, \dots, E_{L_2-1|l_1, m_1, L_2-1}) \triangleq$$

$$\triangleq \min_{l_2 = \overline{1, L_2 - 1}} E_{l_2|l_1, m_1, L_2}^*. \quad (18.d)$$

The following theorem is an analog of Theorem 1 [1] for the dependent object.

Theorem 2: If for given \mathbf{x}_1, m_1, l_1 , with $m_1, l_1 = \overline{1, L_1}$, all conditional PDs G_{l_2/l_1} , $l_2 = \overline{1, L_2}$, are different in the sense that $D(G_{l_2/l_1} \| G_{m_2/l_1} | Q_{\mathbf{x}_1}) > 0$, $l_2 \neq m_2$, $m_2 = \overline{1, L_2}$, when the positive numbers $E_{1/l_1, m_1, 1}, E_{2/l_1, m_1, 2}, \dots, E_{L_2-1/l_1, m_1, L_2-1}$ are such that the following inequalities hold

$$E_{1/l_1, m_1, 1} < \min_{l_2=2, \dots, L_2} D(G_{l_2/l_1} \| G_{1/l_1} | Q_{\mathbf{x}_1}), \quad (19)$$

and for $m_2 = \overline{2, L_2-1}$,

$$E_{m_2/l_1, m_1, m_2} < \min_{l_2=m_2+1, \dots, L_2} D(G_{l_2/l_1} \| G_{m_2/l_1} | Q_{\mathbf{x}_1}), \quad \min_{l_2=1, m_2-1} E_{l_2/l_1, m_1, m_2}^*(E_{l_2/l_1, m_1, l_2}),$$

then there exists a LAO sequence of tests, the reliability matrix of which $E(\varphi_2^*) = E^*(l_1, m_1, \mathbf{x}_1) = \{E_{l_2/l_1, m_1, m_2}^*\}$ is defined in (18) and all elements of it are positive.

When one of the inequalities (19) is violated, then at least one element of the matrix $E^*(l_1, m_1, \mathbf{x}_1)$ is equal to 0.

Proof: For $\mathbf{x}_1 \in \mathcal{X}^N$, $\mathbf{x}_2 \in T_{Q,V}^N(X_2 | \mathbf{x}_1)$ the conditional probability $G_{m_2/l_1}^N(\mathbf{x}_2 | \mathbf{x}_1)$ can be presented as follows (for brevity in this proof we will write $V(V \in \mathcal{V}_N(\mathcal{X}, Q_{\mathbf{x}_1}))$ instead of $V_{\mathbf{x}_1, \mathbf{x}_2}$ and $Q(Q \in \mathcal{P}_N(\mathcal{X}))$ instead of $Q_{\mathbf{x}_1}$

$$\begin{aligned} G_{m_2/l_1}^N(\mathbf{x}_2 | \mathbf{x}_1) &= \prod_{n=1}^N G_{m_2/l_1}(x_n^2 | x_n^1) = \prod_{x^1, x^2} G_{m_2/l_1}(x^2 | x^1)^{N(x^1, x^2 | \mathbf{x}_1, \mathbf{x}_2)} = \\ &= \prod_{x^1, x^2} G_{m_2/l_1}(x^2 | x^1)^{NQ(x^1)V(x^2|x^1)} = \exp\{N \sum_{x^1, x^2} (-Q(x^1)V(x^2|x^1) \log \frac{V(x^2|x^1)}{G_{m_2/l_1}(x^2|x^1)} + \\ &+ Q(x^1)V(x^2|x^1) \log V(x^2|x^1))\} = \exp\{-N[D(V \| G_{m_2/l_1} | Q) + H_{Q,V}(X_2 | X_1)]\}. \end{aligned} \quad (20)$$

We will prove that the sequence of tests φ_2^* , defined for each $\mathbf{x}_1 \in \mathcal{X}^N$ and l_1, m_1 by the following assembly of sets of types

$$\mathcal{B}_{l_2/l_1}^{(N)}(\mathbf{x}_1) = \bigcup_{V \in \mathcal{R}_{l_2/l_1}(Q)} T_{Q,V}^N(X_2 | \mathbf{x}_1), \quad l_2 = \overline{1, L_2}, \quad (21)$$

is LAO and its matrix $E^*(l_1, m_1, \mathbf{x}_1)$ is defined in (18). First we show that each N -vector \mathbf{x}_2 is in one and only in one of $\mathcal{B}_{l_2/l_1}^{(N)}(\mathbf{x}_1)$, that is

$$\mathcal{B}_{l_2/l_1}^{(N)}(\mathbf{x}_1) \cap \mathcal{B}_{m_2/l_1}^{(N)}(\mathbf{x}_1) = \emptyset, \quad l_2 = \overline{1, L_2-1}, \quad m_2 = \overline{l_2+1, L_2} \quad \text{and} \quad \bigcup_{l_2=1}^{L_2} \mathcal{B}_{l_2/l_1}^{(N)}(\mathbf{x}_1) = \mathcal{X}^N.$$

Really, (17.b) and (21) show that

$$\mathcal{B}_{l_2/l_1}^{(N)}(\mathbf{x}_1) \cap \mathcal{B}_{l_2+1/l_1}^{(N)}(\mathbf{x}_1) = \emptyset, \quad l_2 = \overline{1, L_2-1},$$

and for $l_2 = \overline{1, L_2-2}$, $m_2 = \overline{l_2+1, L_2-1}$, for each $\mathbf{x}_1 \in \mathcal{X}^N$ let us consider arbitrary $\mathbf{x}_2 \in \mathcal{B}_{l_2/l_1}^{(N)}(\mathbf{x}_1)$. It follows from (17.a) and (21) that if $Q \in \mathcal{P}_N(\mathcal{X})$ there are $V \in \mathcal{V}_N(\mathcal{X}, Q)$ such that $D(V \| G_{l_2/l_1} | Q) \leq E_{l_2/l_1, m_1, l_2}$ and $\mathbf{x}_2 \in T_{Q,V}^N(X_2 | \mathbf{x}_1)$. From (17) - (19) we have $E_{m_2/l_1, m_1, m_2} < E_{l_2/l_1, m_1, m_2}^*(E_{l_2/l_1, m_1, l_2}) < D(V \| G_{m_2/l_1} | Q)$. From definition (21) we see that $\mathbf{x}_2 \notin \mathcal{B}_{m_2/l_1}^{(N)}(\mathbf{x}_1)$.

Now for $m_2 = \overline{1, L_2 - 1}$, using (5), (16), (17), (18), (20) and (21) we can upper estimate $\alpha_{m_2|l_1, m_1, m_2}^{*N}$ as follows:

$$\begin{aligned} \alpha_{m_2|l_1, m_1, m_2}^{*N} &= G_{m_2/l_1}^N(\overline{B_{m_2/l_1}^{(N)}}(\mathbf{x}_1)|\mathbf{x}_1) = \\ &= G_{m_2/l_1}^N\left(\bigcup_{V: D(V||G_{m_2/l_1}|Q) > E_{m_2|l_1, m_1, m_2}} T_{Q,V}^N(X_2|\mathbf{x}_1)|\mathbf{x}_1\right) \leq \\ &\leq (N+1)^{|\mathcal{X}|^2} \sup_{V: D(V||G_{m_2/l_1}|Q_{\mathbf{x}_1}) > E_{m_2|l_1, m_1, m_2}} G_{m_2/l_1}(T_{Q,V}^N(X_2|\mathbf{x}_1)|\mathbf{x}_1) \leq \\ &\leq (N+1)^{|\mathcal{X}|^2} \sup_{V: D(V||G_{m_2/l_1}|Q) > E_{m_2|l_1, m_1, m_2}} \exp\{-ND(V||G_{m_2/l_1}|Q)\} \leq \\ &\leq \exp\{-N[\inf_{V: D(V||G_{m_2/l_1}|Q) > E_{m_2|l_1, m_1, m_2}} D(V||G_{m_2/l_1}|Q) - o_N(1)]\} \leq \\ &\leq \exp\{-N[E_{m_2|l_1, m_1, m_2} - o_N(1)]\}, \end{aligned}$$

where $o_N(1) \rightarrow 0$ with $N \rightarrow \infty$.

Then we obtain upper and lower estimates for error probabilities from where we can conclude that the elements of reliability matrix of the test φ_2^* are defined in (18):

$$\begin{aligned} \alpha_{l_2|l_1, m_1, m_2}^{*N} &= G_{m_2/l_1}^N(B_{l_2/l_1}^{(N)}(\mathbf{x}_1)|\mathbf{x}_1) = G_{m_2/l_1}^N\left(\bigcup_{V: V \in \mathcal{R}_{l_2/l_1}(Q)} T_{Q,V}^N(X_2|\mathbf{x}_1)|\mathbf{x}_1\right) \leq \\ &\leq (N+1)^{|\mathcal{X}|^2} \sup_{V: V \in \mathcal{R}_{l_2/l_1}(Q)} G_{m_2/l_1}(T_{Q,V}^N(X_2|\mathbf{x}_1)|\mathbf{x}_1) \leq \\ &\leq (N+1)^{|\mathcal{X}|^2} \sup_{V: V \in \mathcal{R}_{l_2/l_1}(Q)} \exp\{-ND(V||G_{m_2/l_1}|Q)\} = \\ &= \exp\{-N[\inf_{V: V \in \mathcal{R}_{l_2/l_1}(Q)} D(V||G_{m_2/l_1}|Q) - o_N(1)]\}, \end{aligned} \quad (22)$$

$$\begin{aligned} \alpha_{l_2|l_1, m_1, m_2}^{*N} &= G_{m_2/l_1}^N(B_{l_2/l_1}^{(N)}(\mathbf{x}_1)|\mathbf{x}_1) = G_{m_2/l_1}^N\left(\bigcup_{V: V \in \mathcal{R}_{l_2/l_1}(Q)} T_{Q,V}^N(X_2|\mathbf{x}_1)|\mathbf{x}_1\right) \geq \\ &\geq \sup_{V: V \in \mathcal{R}_{l_2/l_1}(Q)} G_{m_2/l_1}(T_{Q,V}^N(X_2|\mathbf{x}_1)|\mathbf{x}_1) \geq \\ &\geq (N+1)^{-|\mathcal{X}|^2} \sup_{V: V \in \mathcal{R}_{l_2/l_1}(Q)} \exp\{-ND(V||G_{m_2/l_1}|Q)\} = \\ &= \exp\{-N[\inf_{V: V \in \mathcal{R}_{l_2/l_1}(Q)} D(V||G_{m_2/l_1}|Q) + o_N(1)]\}. \end{aligned} \quad (23)$$

Taking into account (22), (23) and the continuity of the functional $D(V||G_{m_2/l_1}|Q)$ we obtain that $\lim_{N \rightarrow \infty} -N^{-1} \log \alpha_{l_2|l_1, m_1, m_2}^{*N}$ exists and in correspondence with (18b) equals to $E_{l_2|l_1, m_1, m_2}^*$.

Thus $E_{l_2|l_1, m_1, m_2}(\varphi_2^*) = E_{l_2|l_1, m_1, m_2}^*$, $m_2 = \overline{1, L_2}$, $l_2 = \overline{1, L_2}$.

The proof of the first part of the theorem will be finished if we show that the sequence of the tests φ_2^* is LAO, that is for given $E_{1|l_1, m_1, 1}, \dots, E_{L_2-1|l_1, m_1, L_2-1}$ and for any sequence of tests φ_2^{**} for all $m_2, l_2 = \overline{1, L_2}$, $E_{l_2|l_1, m_1, m_2}^{**} \leq E_{l_2|l_1, m_1, m_2}^*$.

Consider sequence φ_2^{**} of tests, which is defined by the sets $\mathcal{D}_{1/l_1}^{(N)}, \mathcal{D}_{2/l_1}^{(N)}, \dots, \mathcal{D}_{L_2/l_1}^{(N)}$ such that $E_{l_2|l_1, m_1, m_2}^{**} \geq E_{l_2|l_1, m_1, m_2}^*$ for some l_2, m_2 . For large enough N we can replace this condition with the following inequality

$$\alpha_{l_2|l_1, m_1, m_2}^{**N} \leq \alpha_{l_2|l_1, m_1, m_2}^{*N}. \quad (24)$$

Let us examine the sets $\mathcal{D}_{l_2/l_1}^{(N)} \cap \mathcal{B}_{l_2/l_1}^{(N)}$, $l_2 = \overline{1, L_2 - 1}$. This intersection cannot be empty, because in that case

$$\begin{aligned} \alpha_{l_2/l_1, m_2}^{**N} &= G_{l_2/l_1}^N(\overline{\mathcal{D}}_{l_2/l_1}^{(N)} | \mathbf{x}_1) \geq G_{l_2/l_1}^N(\mathcal{B}_{l_2/l_1}^{(N)} | \mathbf{x}_1) \geq \\ &\geq \sup_{V: D(V|G_{l_2/l_1}|Q) \leq E_{l_2/l_1, m_2, l_2}} G_{l_2/l_1}^*(\mathcal{T}_{V,Q}^N(X_2 | \mathbf{x}_1)) \geq \exp\{-N(E_{l_2/l_1, m_2, l_2} + o_N(1))\}, \end{aligned}$$

which is at variance with (24). Let us show that $\mathcal{D}_{l_2/l_1}^{(N)} \cap \mathcal{B}_{m_2/l_1}^{(N)} = \emptyset$, $m_2, l_2 = \overline{1, L_2 - 1}$, $l_2 \neq m_2$. If there exists V such that $D(V|G_{m_2/l_1}|Q) \leq E_{m_2/l_1, m_1, m_2}$ and $\mathcal{T}_{V,Q}^N(X_2 | \mathbf{x}_1) \in \mathcal{D}_{l_2/l_1}^{(N)}$, then

$$\alpha_{l_2/l_1, m_1, m_2}^{**N} = G_{m_2/l_1}^N(\mathcal{D}_{l_2/l_1}^{(N)} | \mathbf{x}_1) > G_{m_2/l_1}^N(\mathcal{T}_{V,Q}^N(X_2 | \mathbf{x}_1)) \geq \exp\{-N(E_{m_2/l_1, m_1, m_2} + o_N(1))\}.$$

When $\emptyset \neq \mathcal{D}_{l_2/l_1}^{(N)} \cap \mathcal{T}_{V,Q}^N(X_2 | \mathbf{x}_1) \neq \mathcal{T}_{V,Q}^N(X_2 | \mathbf{x}_1)$, we also obtain that

$$\begin{aligned} \alpha_{l_2/l_1, m_1, m_2}^{**N} &= G_{m_2/l_1}^N(\mathcal{D}_{l_2/l_1}^{(N)} | \mathbf{x}_1) > G_{m_2/l_1}^N(\mathcal{D}_{l_2/l_1}^{(N)} \cap \mathcal{T}_{V,Q}^N(X_2 | \mathbf{x}_1) | \mathbf{x}_1) \geq \\ &\geq \exp\{-N(E_{m_2/l_1, m_2, m_2} + o_N(1))\}. \end{aligned}$$

Thus we conclude that $E_{l_2/l_1, m_1, m_2}^{**} < E_{m_2/l_1, m_1, m_2}$, which contradicts to (7). Hence we obtain that $\mathcal{D}_{l_2/l_1}^{(N)} \cap \mathcal{B}_{l_2/l_1}^{(N)} = \mathcal{B}_{l_2/l_1}^{(N)}$ for $l_2 = \overline{1, L_2 - 1}$. The following intersection $\mathcal{D}_{l_2/l_1}^{(N)} \cap \mathcal{B}_{L_2/l_1}^{(N)}$ is empty too, because otherwise

$$\alpha_{L_2/l_1, m_1, m_2}^{**N} \geq \alpha_{L_2/l_1, m_1, m_2}^{*N},$$

which contradicts to (24), it means that $\mathcal{D}_{l_2/l_1}^{(N)} = \mathcal{B}_{l_2/l_1}^{(N)}$, for all $l_2 = \overline{1, L_2}$. So we have proved that test φ_2^* is the unique LAO test.

The proof of the second assertion of the theorem is simple. If one of the conditions (19) is violated, then it follows from (17) and (18) that at least one of the elements $E_{l_2/l_1, m_1, m_2}$ is equal to 0.

3 LAO Testing of L_1, L_2 Hypotheses for Two Dependent Objects

Let us define the following subsets of $\mathcal{P}(\mathcal{X})$ for given strictly positive elements $E_{L_1, l_2/l_1, l_2}$, $E_{l_1, L_2/l_1, l_2}$, $l_1 = \overline{1, L_1 - 1}$, $l_2 = \overline{1, L_2 - 1}$:

$$\begin{aligned} \mathcal{R}_{l_1} &\triangleq \{Q : D(Q|G_{l_1}) \leq E_{L_1, l_2/l_1, l_2}\}, \quad l_1 = \overline{1, L_1 - 1}, \quad l_2 = \overline{1, L_2 - 1}, \\ \mathcal{R}_{l_2/l_1}(Q_{\mathbf{x}_1}) &\triangleq \{V : D(V|G_{l_2/l_1}|Q_{\mathbf{x}_1}) \leq E_{l_1, L_2/l_1, l_2}\}, \quad l_1 = \overline{1, L_1 - 1}, \quad l_2 = \overline{1, L_2 - 1}, \\ \mathcal{R}_{L_1} &\triangleq \{Q : D(Q|G_{l_1}) > E_{L_1, l_2/l_1, l_2}\}, \quad l_1 = \overline{1, L_1 - 1}, \quad l_2 = \overline{1, L_2 - 1}, \\ \mathcal{R}_{L_2/l_1}(Q_{\mathbf{x}_1}) &\triangleq \{V : D(V|G_{L_2/l_1}|Q_{\mathbf{x}_1}) > E_{l_1, L_2/l_1, l_2}\}, \quad l_1 = \overline{1, L_1 - 1}, \quad l_2 = \overline{1, L_2 - 1}. \end{aligned}$$

Assume also that

$$E_{l_1, L_2/l_1, l_2}^* \triangleq E_{l_1, L_2/l_1, l_2}, \quad E_{L_1, l_2/l_1, l_2}^* \triangleq E_{L_1, l_2/l_1, l_2}, \quad l_1 = \overline{1, L_1 - 1}, \quad l_2 = \overline{1, L_2 - 1}, \quad (25.a)$$

$$E_{l_1, l_2 | m_1, l_2}^* \triangleq \inf_{Q \in R_{l_1, l_2}} D(Q \| G_{m_1}), \quad m_1 \neq l_1 \quad (25.b)$$

$$E_{l_1, l_2 | l_1, m_2}^* \triangleq \inf_{V: V \in R_{l_2/l_1}(Q_{x_1})} D(V \| G_{m_2/m_1} | Q_{x_1}), \quad m_2 \neq l_2 \quad (25.c)$$

$$E_{l_1, l_2 | m_1, m_2}^* \triangleq E_{m_1, l_2 | m_1, m_2}^* + E_{l_1, m_2 | m_1, m_2}^*, \quad m_1 \neq l_1, \quad i = 1, 2, \quad (25.d)$$

$$E_{m_1, m_2 | m_1, m_2}^* \triangleq \min_{(l_1, l_2) \neq (m_1, m_2)} E_{l_1, l_2 | m_1, m_2}^* \quad (25.e)$$

Theorem 3: If all distributions G_{m_1} , $m_1 = \overline{1, L_1}$, are different, that is $D(G_{l_1} \| G_{m_1}) > 0$, $l_1 \neq m_1$, $l_1, m_1 = \overline{1, L_1}$, and all conditional distributions G_{l_2/l_1} , $l_2 = \overline{1, L_2}$, are also different for all $l_1 = \overline{1, L_1}$, in the sense that $D(G_{l_2/l_1} \| G_{m_2/l_1} | Q_{x_1}) > 0$, $l_2 \neq m_2$, then the following two statements are valid.

When given elements $E_{L_1, l_2 | l_1, l_2}$ and $E_{l_1, L_2 | l_1, l_2}$, $l_1 = \overline{1, L_1 - 1}$, $l_2 = \overline{1, L_2 - 1}$, meet the following conditions

$$0 < E_{L_1, l_2 | l_1, l_2} < \min_{m_1 = \overline{2, L_1}} D(G_{m_1} \| G_1), \quad (26.a)$$

$$0 < E_{l_1, L_2 | l_1, l_2} < \min_{m_2 = \overline{2, L_2}} D(G_{m_2/l_1} \| G_{l_1/l_1} | Q_{x_1}), \quad (26.b)$$

$$0 < E_{L_1, l_2 | l_1, l_2} < \min \left[\min_{m_1 = \overline{1, l_1 - 1}} E_{m_1, l_2 | l_1, l_2}^*, \min_{m_1 = l_1 + 1, L_1} D(G_{m_1} \| G_{l_1}) \right], \quad l_1 = \overline{2, L_1 - 1}, \quad (26.c)$$

$$0 < E_{l_1, L_2 | l_1, l_2} < \min \left[\min_{m_2 = \overline{1, l_2 - 1}} E_{l_1, m_2 | l_1, l_2}^*, \min_{m_2 = l_2 + 1, L_2} D(G_{m_2/l_1} \| G_{l_2/l_1} | Q_{x_1}) \right], \quad l_2 = \overline{2, L_2 - 1}, \quad (26.d)$$

then there exists a LAO test sequence Φ^* , the reliability matrix of which $E(\Phi^*) = \{E_{l_1, l_2 | m_1, m_2}(\Phi^*)\}$ is defined in (25) and all elements of it are positive.

When even one of the inequalities (26) is violated, then at least one element of the matrix $E(\Phi^*)$ is equal to 0.

Proof: It is known from [5] that $E_{l_1, l_1} = E_{L_1, l_1}$, $l_1 = \overline{1, L_1 - 1}$. Analogously we can deduce that

$$E_{l_2, l_2 | m_1, l_2}^* = E_{L_2, l_2 | m_1, l_2}^*, \quad l_2 = \overline{1, L_2 - 1}. \quad (27)$$

Applying theorem of Kuhn-Tucker in (18.b) we can show that the elements $E_{l_2, l_2 | m_1, l_2}^*$, $l_2 = \overline{1, L_2 - 1}$ can be determined by elements E_{l_2, l_1, m_1, m_2}^* , $m_2 \neq l_2$, $l_2 = \overline{1, L_2}$,

$$E_{l_2, l_1, m_1, l_2}^* (E_{l_2, l_1, m_1, m_2}^*) \triangleq \inf_{V: D(V \| G_{m_2/l_1} | Q_{x_1}) \leq E_{l_2, l_1, m_1, m_2}^*} D(V \| G_{l_2/l_1} | Q_{x_1}).$$

From (19) it is clear that E_{m_2, l_1, m_1, m_2}^* can be equal only to one of E_{l_2, l_1, m_1, m_2}^* , $l_2 = \overline{m_2 + 1, L_2}$. Assume that (27) is not correct, that is $E_{m_2, l_1, m_1, m_2}^* = E_{l_2, l_1, m_1, m_2}^*$, $l_2 = \overline{m_2 + 1, L_2 - 1}$. From (18.b) it follows that

$$\begin{aligned} E_{l_2, l_1, m_1, l_2}^* (E_{l_2, l_1, m_1, m_2}^*) &\triangleq \inf_{V: D(V \| G_{m_2/l_1} | Q_{x_1}) \leq E_{l_2, l_1, m_1, m_2}^*} D(V \| G_{l_2/l_1} | Q_{x_1}) = \\ &= \inf_{V: D(V \| G_{m_2/l_1} | Q_{x_1}) \leq E_{m_2, l_1, m_1, m_2}^*} D(V \| G_{l_2/l_1} | Q_{x_1}) = E_{m_2, l_1, m_1, l_2}^*, \\ &m_2, l_2 = \overline{1, L_2 - 1}, m_2 < l_2, \end{aligned}$$

but from conditions (19) it follows that $E_{l_2/l_1, m_1, l_2}^* < E_{m_2/l_1, m_1, l_2}^*$ for $m_2 = \overline{1, l_2 - 1}$. Our assumption is not true, thus (27) is valid. Hence we can rewrite the inequalities (15) and (19) as follows:

$$0 < E_{l_1/l_1} < \min_{m_1 = \overline{2, L_1}} D(G_{m_1} || G_1), \quad (28.a)$$

$$0 < E_{L_2/l_1, m_1, l_1} < \min_{m_2 = \overline{2, L_2}} D(G_{m_2/l_1} || G_{l_1/l_1} | Q_{x_1}), \quad (28.b)$$

$$0 < E_{L_1/l_1} < \min \left[\min_{m_1 = \overline{1, l_1 - 1}} E_{m_1/l_1}^*, \min_{m_1 = \overline{l_1 + 1, L_1}} D(G_{m_1} || G_{l_1/l_1}) \right], \quad l_1 = \overline{2, L_1 - 1}, \quad (28.c)$$

$$0 < E_{L_2/l_1, m_1, l_2} < \min \left[\min_{m_2 = \overline{1, l_2 - 1}} E_{m_2/l_1, m_1, l_2}^*, \min_{m_2 = \overline{l_2 + 1, L_2}} D(G_{m_2/l_1} || G_{l_2/l_1} | Q_{x_1}) \right], \quad l_2 = \overline{2, L_2 - 1}, \quad (28.d)$$

According to Theorem 1 and Theorem 2 there exist LAO sequences of tests φ_1^* and φ_2^* , for the first and the second objects, such that the elements of the matrices E^* and $E^*(l_1, m_1, x_1)$ are determined in (14) and (18). The inequalities (28.a), (28.c) are equivalent to the inequalities (15) and (28.b), (28.d) are equivalent to the inequalities (19). Then using Lemma we deduce that the reliability matrix $E(\Phi^*)$ can be determined in (25). When one of inequalities of (26.a) and (26.c) ((26.b) and (26.d)) is violated then using (11.b) and (25.b) ((11.c) and (25.c)) we see that some elements of matrix must be equal to 0.

Corollary: When the objects are dependent only statistically, that is the L_1 PDs are known for the first objects, and the second object depending on distribution index of the first can be distributed according to one of given conditional distributions, then $G_{m_2/l_1}(x^2 | x^1) = G_{m_2/l_1}(x^2)$, $m_2 = \overline{1, L_2}$, $l_1 = \overline{1, L_1}$. Thus the first object characterized by RV X_1 can have one of given L_1 PDs $G_{l_1} = \{G_{l_1}(x^1), x^1 \in \mathcal{X}\}$, $l_1 = \overline{1, L_1}$, from $\mathcal{P}(\mathcal{X})$ and the second object characterized by RV X_2 can have one of $L_1 \times L_2$ conditional PDs $G_{l_2/l_1} = \{G_{l_2/l_1}(x^2), x^2 \in \mathcal{X}\}$, $l_1 = \overline{1, L_1}$, $l_2 = \overline{1, L_2}$ and it follows that $G_{m_2/l_1}^N(x_2 | x_1) = G_{m_2/l_1}^N(x_2)$, $m_2 = \overline{1, L_2}$, $l_1 = \overline{1, L_1}$. Hence sequence of tests φ_2^* is depended on vector x_2 and on the index of the hypothesis l_1 adopted for X_1 and we obtain the generalization of the case examined in [11].

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Ստոխաստիկորեն կախված օբյեկտների զույգի նկատմամբ վարկածների օպտիմալ ստուգման մասին

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Ամփոփում

Լուծված է ստոխաստիկորեն կախյալ երկու օբյեկտների նկատմամբ վարկածների լոգարիթմորեն ասիմպտոտորեն օպտիմալ ստուգման խնդիրը: Առաջին օբյեկտը կարող է բաշխված լինել տրված L_1 հավանականային բաշխումներից մեկով, իսկ երկրորդը՝ կախված առաջինի բաշխումից և դիտարկվող պահին նրա վիճակից, տրված $L_1 \times L_2$ պայմանական հավանականային բաշխումներից մեկով: Ուսումնասիրվել է օբյեկտների նկատմամբ վարկածների տեստավորման սխեմների հավանականությունների ցուցիչների (հուսալիությունների) փոխկախվածությունների մատրիցը: