

Guessing Subject to Distortion and Reliability Criteria in the Shannon Cipher System

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Abstract

A problem of a discrete memoryless source messages guessing within given distortion level Δ for the Shannon cipher system is solved. The security of this system is measured by the expected number of required guesses of the wiretapper needed for the reconstruction of the source messages on the base of cryptograms. In addition to the problem studied by Merhav and Arikan we demand that for a given guessing list, distortion level $\Delta \geq 0$ and reliability $E > 0$ the probability that distortions between blocklength N source messages and each of all first $L(N)$ guessing vectors will be larger than Δ , does not exceed 2^{-NE} .

For given key rate R' the minimum (over all guessing lists) of the maximal (over all encryption functions) $L(N)$ and of the expected number of required guesses, with respect to distortion and reliability criteria, are found.

1 Introduction

We investigate the wiretapper guessing problem with respect to fidelity and reliability conditions for the Shannon cipher system (see Fig.). Encrypted messages of a discrete memoryless stationary source must be communicated as securely as possible to a legitimate receiver, which has access to a common key-vector which is independent of source messages.

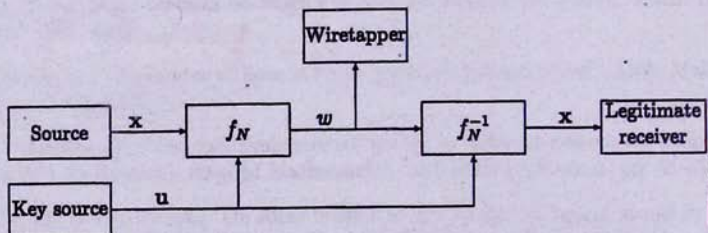


Fig. The Shannon cipher system with a guessing wiretapper.

The key-vector is transmitted by the special secure channel protected against wiretappers. The transmitter encodes the source message and key-vector and sends a cryptogram over a

*The work of the author was supported by INTAS YSF Grant 00-4163.

public channel to a legitimate receiver which based on the cryptogram and key-vector recovers the original message by the inverse, decryption function. The wiretapper that eavesdrop a public channel aims to decrypt the original source message in the framework of given distortion and reliability on the base of cryptogram, source statistic and encryption function. The wiretapper has a testing mechanism by which he can know whether the estimate is successful (is within given distortion level). The problem is to determine the minimum of the maximal (over all possible encryption functions) expected numbers of sequential guesses until the satisfactory message will be found.

The guessing problem was first considered by Massey [20] and later was investigated in [3]-[7], [9], [21]. The guessing problem subject to fidelity criterion was considered by Arikan and Merhav in [5]-[7], for reliability criterion - in [14], for the Shannon cipher system with exact reconstruction - in [21]. The Shannon cipher system with wiretapper reconstructing source messages subject to fidelity criterion were studied by Yamamoto in [23]. We study a combination of these problems with additional reliability criterion: a generalization proposed by E. Haroutunian and B. Mekoush [18] of the usual Shannon rate-distortion concept as a rate-reliability-distortion dependence. The idea is that the minimum of the maximal (over all possible encryption functions) expected number of required guesses is considered in relation to demands of guesser not only to distortion level but also to the error probability exponent (reliability). This approach was investigated for various multiterminal systems (see for example [13]-[19], [22]).

Now we pass to more detailed definitions. The source $\{X\}$ is defined as a sequence $\{X_i\}_{i=1}^{\infty}$ of discrete, independent, identically distributed (d. i. i. d.) random variables (RV) taking values in the finite set \mathcal{X} of messages of the source. Let

$$P^* = \{P^*(x), x \in \mathcal{X}\} \quad (1)$$

be the given probability distribution (PD) of source messages. Since we study the memoryless source,

$$P^{*N}(x) = \prod_{n=1}^N P^*(x_n).$$

The key $\{U\}$ is a sequence $\{U_i\}_{i=1}^{\infty}$ of d. i. i. d. RV taking values in the set $\mathcal{U} = \{0, 1\}$. Let $\mathbf{X} = (X_1, X_2, \dots, X_N)$ be a random sequence of N messages, $\mathbf{U} = (U_1, U_2, \dots, U_K)$ be a sequence of K purely random bits independent of the source messages vector \mathbf{X} . Denote by $\hat{\mathcal{X}}$ the reconstruction by the wiretapper of the source message, with values in finite set $\hat{\mathcal{X}}$, which in general is different from \mathcal{X} and is wiretapper reproduction alphabet. \mathcal{X}^N and $\hat{\mathcal{X}}^N$ denote the N -th order Cartesian powers of the sets \mathcal{X} and $\hat{\mathcal{X}}$, respectively, \mathcal{U}^K - the K -th order Cartesian power of the set \mathcal{U} . We consider a single-letter distortion measure between source and wiretapper reproduction messages:

$$d: \mathcal{X} \times \hat{\mathcal{X}} \rightarrow [0; \infty).$$

The distortion measure between a source vector $\mathbf{x} = (x_1, x_2, \dots, x_N) \in \mathcal{X}^N$ and a wiretapper reproduction vector $\hat{\mathbf{x}} = (\hat{x}_1, \hat{x}_2, \dots, \hat{x}_N) \in \hat{\mathcal{X}}^N$ is defined as average of the components' distortions:

$$d(\mathbf{x}, \hat{\mathbf{x}}) = N^{-1} \sum_{n=1}^N d(x_n, \hat{x}_n).$$

Distribution $P_1^* = \{1/2, 1/2\}$ is the PD of the key bits. The key-vector $\mathbf{u} = (u_1, u_2, \dots, u_K)$ is a sequence of these bits and $P_1^{*K}(\mathbf{u}) = \frac{1}{2^K}$.

Let

$$f_N: \mathcal{X}^N \times \mathcal{U}^K \rightarrow \mathcal{W}(N, K)$$

be an encryption function and $\mathcal{W}(N, K)$ be the set of all corresponding cryptograms. It is assumed that this function is invertible given the key, that is there exists the decryption function

$$f_N^{-1}: \mathcal{W}(N, K) \times \mathcal{U}^K \rightarrow \mathcal{X}^N.$$

For a cryptogram $w = f_N(x, u)$, the ordered list of sequential guesses

$$\mathcal{G}_N(w) = \{\hat{x}_1(w), \hat{x}_2(w), \dots\} \quad (\hat{x}_l(w) \in \hat{\mathcal{X}}^N, l = 1, 2, \dots, |\hat{\mathcal{X}}|^N)$$

is called the guessing strategy of the wiretapper. For given guessing strategy $\mathcal{G}_N(w)$, $w \in \mathcal{W}(N, K)$ we name a guessing function and note $G_N(x | w)$ the function that maps each vector $x \in \mathcal{X}^N$ into a positive integer:

$$G_N: \mathcal{X}^N \times \mathcal{W}(N, K) \rightarrow \{1, 2, 3, \dots, L(N)\},$$

which is the index l of the first such guessing vector $\hat{x}_l(w) \in \mathcal{G}_N(w)$ that $d(x, \hat{x}_l(w)) \leq \Delta$. In other words, l is the number of sequential guesses of the wiretapper for a source vector $x \in \mathcal{X}^N$ until the successful estimate $\hat{x}_l(w) \in \mathcal{G}_N(w)$ will be found. Let $L(N)$ is prescribed limitation of the number of guesses.

For each distortion level $\Delta \geq 0$, positive number $L(N) < |\hat{\mathcal{X}}|^N$ and a guessing strategy $\mathcal{G}_N(w)$ consider the following set of successfully deciphered vectors of messages

$$\mathcal{A}(L(N), \mathcal{G}_N(w), \Delta) = \{x: \exists u, \exists l, l \leq L(N): f(x, u) = w, d(x, \hat{x}_l(w)) \leq \Delta\},$$

and the probability of error

$$e(L(N), \mathcal{G}_N(w), \Delta) = 1 - P^{*N}(\mathcal{A}(L(N), \mathcal{G}_N(w), \Delta)).$$

Below log-s and exp-s are taken to the base 2.

Let R' be the key rate:

$$R' = N^{-1} \log 2^K = K.$$

A pair of guessing rates (R'', R) is called (R', E, Δ) -achievable for $E > 0$, $\Delta \geq 0$ and R' , if for every $\varepsilon > 0$, encryption function f and sufficiently large N there exists a guessing strategy $\mathcal{G}_N(w)$ such that

$$N^{-1} \log L(N) \leq R'' + \varepsilon,$$

$$N^{-1} \log E_{P^*, P^*} \{G_N(X | W)\} \leq R + \varepsilon,$$

and

$$e(L(N), \mathcal{G}_N(w), \Delta) \leq \exp\{-NE\}.$$

Let us denote by $R_G(P^*, R', E, \Delta)$ and call the key-distortion-reliability guessing rates region the set of all (R', E, Δ) -achievable pairs of guessing rates. When $E \rightarrow \infty$, $\mathcal{X} \equiv \hat{\mathcal{X}}$ and $R'' = \log |\mathcal{X}|$, $R_G(P^*, R', E, \Delta)$ becomes the key guessing rate function $R_G(P^*, R')$, studied by Merhav and Arikan in [21].

In the next section we specify the region of key-distortion-reliability guessing rates. The proof is presented in the section 3.

2 Formulation of Result

Let $P = \{P(x), x \in \mathcal{X}\}$ be a PD on \mathcal{X} and $Q = \{Q(\hat{x} | x), x \in \mathcal{X}, \hat{x} \in \hat{\mathcal{X}}\}$ be a conditional PD on $\hat{\mathcal{X}}$ for given x .

Consider for given $E > 0$ the following set of PD P :

$$\alpha(E) = \{P : D(P \| P^*) \leq E\}$$

with divergence

$$D(P \| P^*) = \sum_x P(x) \log \frac{P(x)}{P^*(x)}.$$

Denote by $\Phi(P, \Delta) = Q_P$ a function, which puts into the correspondence to PD P some such conditional PD Q_P that for given Δ the following condition is fulfilled:

$$E_{P, Q_P} d(X, \hat{X}) = \sum_x P(x) Q_P(\hat{x} | x) d(x, \hat{x}) \leq \Delta.$$

Denote by $\mathcal{M}(P, \Delta)$ the set of all such functions $\Phi(P, \Delta)$ for given Δ and P . Below for brevity we shall just write $\Phi(P)$.

We use the following notations for entropy and information:

$$H_P(X) = - \sum_x P(x) \log P(x),$$

$$I_{P, Q_P}(X \wedge \hat{X}) = \sum_{x, \hat{x}} P(x) Q_P(\hat{x} | x) \log \frac{Q_P(\hat{x} | x)}{\sum_x P(x) Q_P(\hat{x} | x)}.$$

We denote by $R(P, \Delta)$ the rate-distortion function for given PD P (see [11]):

$$R(P, \Delta) = \min_{\Phi(P) \in \mathcal{M}(P, \Delta)} I_{P, \Phi(P)}(X \wedge \hat{X}),$$

and by $R(P^*, E, \Delta)$ the rate-reliability-distortion function for PD of source messages P^* (see [18]):

$$R(P^*, E, \Delta) = \max_{P \in \alpha(E)} \min_{\Phi(P) \in \mathcal{M}(P, \Delta)} I_{P, \Phi(P)}(X \wedge \hat{X}).$$

Let us introduce the following region:

$$\widetilde{\mathcal{R}}_G(P^*, R', E, \Delta) = \{(R'', R) :$$

$$R'' \geq \min(R', R(P^*, E, \Delta)), \quad (1)$$

$$R \geq \max_{P \in \alpha(E)} [\min(R', R(P, \Delta)) - D(P \| P^*)]. \quad (2)$$

Theorem: For given PD P^* on \mathcal{X} , key rate $R' \geq 0$ and every $E > 0, \Delta \geq 0$,

$$\mathcal{R}_G(P^*, R', E, \Delta) = \widetilde{\mathcal{R}}_G(P^*, R', E, \Delta).$$

Corollary: When $E \rightarrow \infty$, wiretapper requires exact reconstruction of the source messages and $R'' = \log |\mathcal{X}|$, we arrive to the result of Merhav and Arikan from [21]:

$$\lim_{E \rightarrow \infty, \Delta=0, R''=\log|\mathcal{X}|} R_G(P^*, R', E, \Delta) = \max_P [\min(R', H_P(X)) - D(P \| P^*)].$$

3 Proof of the theorem

In the theorem proof we apply the method of types (see [10]–[12]). We begin with the formulation of base notions and relations.

The type P of a vector $\mathbf{x} \in \mathcal{X}^N$ is a PD $P = \{P(x) = N(x|\mathbf{x})/N, x \in \mathcal{X}\}$, where $N(x|\mathbf{x})$ is the number of repetitions of symbol x among x_1, \dots, x_N . The set of all PD P on \mathcal{X} , which for given N are types of vectors in \mathcal{X}^N is denoted by $\mathcal{P}(\mathcal{X}, N)$. The set of vectors \mathbf{x} of type P is denoted by $T_P(\mathcal{X})$ and also called type.

The conditional type of $\hat{\mathbf{x}}$ for given \mathbf{x} is PD $Q_P = \{Q_P(\hat{x}|\mathbf{x}), x \in \mathcal{X}, \hat{x} \in \hat{\mathcal{X}}\}$ if $N(x, \hat{x}|\mathbf{x}, \hat{\mathbf{x}}) = N(x|\mathbf{x})Q_P(\hat{x}|\mathbf{x})$ for all $x \in \mathcal{X}, \hat{x} \in \hat{\mathcal{X}}$. The set of all vectors $\hat{\mathbf{x}} \in \hat{\mathcal{X}}^N$ of conditional type Q_P for given $\mathbf{x} \in T_P(\mathcal{X})$ is denoted by $T_{P,Q_P}(\hat{\mathcal{X}}|\mathbf{x})$. The set of all possible $T_{P,Q_P}(\hat{\mathcal{X}}|\mathbf{x})$ for \mathbf{x} of type P is denoted by $\mathcal{Q}_P(\hat{\mathcal{X}}, P, N)$.

We will use the following useful properties of types ([10]–[12]):

$$|\mathcal{P}_N(\mathcal{X})| \leq (N+1)^{|\mathcal{X}|}, \quad (3)$$

$$|\mathcal{Q}_P(\hat{\mathcal{X}}, P, N)| \leq (N+1)^{|\mathcal{X}||\hat{\mathcal{X}}|}, \quad (4)$$

for any type $P \in \mathcal{P}(\mathcal{X}, N)$

$$(N+1)^{-|\mathcal{X}|} \exp\{NH_P(\mathbf{x})\} \leq |T_P(\mathbf{x})| \leq \exp\{NH_P(\mathbf{x})\}, \quad (5)$$

for any conditional type Q_P and $\mathbf{x} \in T_P(\mathcal{X})$

$$(N+1)^{-|\mathcal{X}||\hat{\mathcal{X}}|} \exp\{NH_{P,Q_P}(\hat{\mathbf{x}}|\mathbf{x})\} \leq |T_{P,Q_P}(\hat{\mathbf{x}}|\mathbf{x})| \leq \exp\{NH_{P,Q_P}(\hat{\mathbf{x}}|\mathbf{x})\}, \quad (6)$$

if $\mathbf{x} \in T_P(\mathcal{X})$, then for any PD Q on $\hat{\mathcal{X}}$

$$Q^N(\mathbf{x}) = \exp\{-N(H_P(\mathbf{x}) + D(P \| Q))\}, \quad (7)$$

$$(N+1)^{-|\mathcal{X}|} \exp\{-ND(P \| Q)\} \leq Q^N(T_P(\mathcal{X})) \leq \exp\{-ND(P \| Q)\}. \quad (8)$$

The proof of the inequality

$$\tilde{\mathcal{R}}_G(P^*, R', E, \Delta) \subseteq \mathcal{R}_G(P^*, R', E, \Delta) \quad (9)$$

is based on the following random coding lemma about covering of types of vectors, which is a modification of the covering lemmas from [1], [2], [11], [19]:

Lemma: Let for $\varepsilon > 0$,

$$L(P, Q) = \exp\{N(I_{P,Q}(X \wedge \hat{X}) + \varepsilon)\}.$$

Then for every type P and conditional distribution Q for N large enough there exist the collections of vectors

$$\{\hat{\mathbf{x}}_l \in T_{P,Q}(\hat{\mathcal{X}}), l = \overline{1, L(P, Q)}\},$$

such that the family of conditional types

$$\{T_{P,Q}(X | \hat{\mathbf{x}}_l), l = \overline{1, L(P, Q)}\}$$

cover $T_P(\mathcal{X})$.

The proof of lemma is similar to the proof of lemmas in [1], [2], [11], [13], [19].

Let us represent \mathcal{X}^N as a family of disjoint types

$$\mathcal{X}^N = \bigcup_{P \in \mathcal{P}(\mathcal{X}, N)} \mathcal{T}_P(X).$$

In the sequel without mention we consider only PD P which are types P for N , that is $P \in \mathcal{P}(\mathcal{X}, N)$.

Let some $\delta > 0$ be given. Then for N large enough the probability of appearance of the source sequences of types beyond $\alpha(E + \delta)$ can be estimated as follows:

$$\begin{aligned} P^{*N} \left(\bigcup_{P \notin \alpha(E+\delta)} \mathcal{T}_P(X) \right) &= \sum_{P \notin \alpha(E+\delta)} P^{*N}(\mathcal{T}_P(X)) \leq \\ &\leq (N+1)^{|\mathcal{X}|} \exp\{-N \min_{P \notin \alpha(E+\delta)} D(P \| P^*)\} \leq \end{aligned}$$

$$\leq \exp\{-NE - N\delta + |\mathcal{X}| \log(N+1)\} \leq \exp\{-N(E + \delta/2)\} \leq \exp\{-NE\}.$$

Here the first inequality follows from (upper estimate in (5)) and the second inequality holds due to the definition of the set $\alpha(E)$. Consequently, in order to obtain the desired level of probability $e(L(N), \mathcal{G}_N(w), \Delta)$, it is sufficient to construct the guessing strategy only for the vectors with types P from $\alpha(E + \delta)$.

First we construct a guessing strategy that ignores the cryptogram. It is possible to order the types P from $\alpha(E + \delta)$ as $\{P_1, P_2, \dots\}$ according to increasing value of corresponding rate-distortion functions $R(P_i, \Delta) = R(i, \Delta)$. (for simplicity of formulae writing instead of P_i we shall write only i):

$$R(i, \Delta) \leq R(i+1, \Delta), \quad 1 \leq i \leq |\alpha(E + \delta)|.$$

For fixed i let the set $\{\hat{x}_{i,l} \in \mathcal{T}_{i, Q_i^{\min}}(\hat{X}), l = \overline{1, L(i, Q_i^{\min})}\}$, with

$$L(i, Q_i^{\min}) = \exp\{N(\min_{\Phi(i) \in \mathcal{M}(i, \Delta)} I_{i, \Phi(i)}(X \wedge \hat{X}) + \varepsilon)\} = \exp\{N(R(i, \Delta) + \varepsilon)\}$$

be a collection of vectors such that according to the lemma for N large enough the set

$$\{\mathcal{T}_{i, Q_i^{\min}}(X | \hat{x}_{i,l}), l = \overline{1, L(i, Q_i^{\min})}\},$$

covers $\mathcal{T}_i(X)$.

Consider the following guessing strategy:

$$\mathcal{G}_N^*(w) = \{\{\hat{x}_{1,l}, l = \overline{1, L(Q_1^{\min})}\}, \{\hat{x}_{2,l}, l = \overline{1, L(Q_2^{\min})}\}, \dots\}.$$

The number of required guesses $G_N^*(\mathbf{x} | w)$ for $\mathbf{x} \in \mathcal{T}_i(X)$, $P_i \in \alpha(E + \delta)$ and any w is upperbounded by

$$G_N^*(\mathbf{x} | w) \leq \sum_{j, j \leq i} L(j, Q_j^{\min}) \leq L(i, Q_i^{\min}) \exp\{N\varepsilon\} = \exp\{N(R(i, \Delta) + 2\varepsilon)\}.$$

Consider now a guessing strategy that using cryptogram w carries out an key-search attack:

$$\mathcal{G}_N^*(w) = \{f_N^{-1}(w, u_1), f_N^{-1}(w, u_2), \dots\},$$

where u_1, u_2, \dots is a sequence of all possible key-vectors of length K . This sequence contains 2^K vectors, therefore for any cryptogram w the number of required guesses $G_N^{**}(x | w)$ for $x \in T_P(X)$, $P \in \alpha(E + \delta)$, is upperbounded by

$$G_N^{**}(x | w) \leq \exp\{K\}.$$

Finally, let us consider the following combined guessing list:

$$G_N^{***}(w) = \{\{\hat{x}_{1l}, l = \overline{1, L(1, Q_1^{\min})}\}, f_N^{-1}(w, u_1), \{\hat{x}_{2l}, l = \overline{1, L(2, Q_2^{\min})}\}, f_N^{-1}(w, u_2), \dots\}.$$

For the given cryptogram w the number of sequential wiretapper guesses for the source vector $x \in T_i(X)$, $P_i \in \alpha(E + \delta)$ is upperbounded by

$$G_N^{***}(x | w) \leq 2 \cdot \min\{\exp\{K\}, \exp\{N(R(i, \Delta) + 2\varepsilon)\} = \exp\{N(\min(R', R(i, \Delta)) + 3\varepsilon)\}.$$

Hence

$$L(N) = \exp\{N(\min(R', R(P^*, E + \delta, \Delta)) + 3\varepsilon)\}.$$

Taking into account independence of key-vectors and source message vectors and using upper estimates (5) for probabilities of the sets $T_i(X)$, $1 \leq i \leq |\alpha(E + \delta)|$, we obtain that

$$\begin{aligned} E_{P^*, P_i^*} \{G_N^{***}(X | W)\} &= \sum_{i: P_i \in \alpha(E + \delta)} \sum_{x \in T_i(X)} \sum_{u \in U^K} P^{*N}(x) P_1^{*K}(u) G_N^{***}(x | f_N(x, u)) \leq \\ &\leq \sum_{i: P_i \in \alpha(E + \delta)} \sum_{x \in T_i(X)} P^{*N}(x) \exp\{N(\min(R', R(i, \Delta)) + 3\varepsilon)\} \sum_{u \in U^K} P_1^{*K}(u) = \\ &= \sum_{i: P_i \in \alpha(E + \delta)} \exp\{N(\min(R', R(i, \Delta)) + 3\varepsilon)\} \sum_{x \in T_i(X)} P^{*N}(x) = \\ &= \sum_{i: P_i \in \alpha(E + \delta)} \exp\{N(\min(R', R(i, \Delta)) + 3\varepsilon)\} P^{*N}(T_i(X)) \leq \\ &\leq \sum_{P \in \alpha(E + \delta)} \exp\{N(-D(P \| P^*) + \min(R', R(P, \Delta)) + 3\varepsilon)\} \leq \\ &\leq \max_{P \in \alpha(E + \delta)} \exp\{N(-D(P \| P^*) + \min(R', R(P, \Delta)) + 4\varepsilon)\} = \\ &= \exp\{N(\max_{P \in \alpha(E + \delta)} (-D(P \| P^*) + \min(R', R(P, \Delta)) + 4\varepsilon))\}. \end{aligned}$$

Therefore a pair of guessing rates (R'', R') such that

$$R'' \geq \frac{1}{N} \log L(N) - \varepsilon \geq \min(R', R(P^*, E + \delta, \Delta)) + 2\varepsilon,$$

$$R \geq \max_{P \in \alpha(E + \delta)} (-D(P \| P^*) + \min(R', R(P, \Delta)) + 4\varepsilon) - \varepsilon$$

is (R', E, Δ) -achievable.

Taking into account the arbitrariness of ε and δ , the continuity of all functions with respect to E , we obtain (9).

Now we shall prove the inclusion

$$\mathcal{R}_G(P^*, R', E, \Delta) \subseteq \widetilde{\mathcal{R}}_G^*(P^*, R', E, \Delta). \quad (10)$$

To obtain lower bounds on $L(N)$ and $E_{P^*, P_1^*}\{G_N^*(X | W)\}$ we may assume that the guesser is informed of the type $T_P(X)$ of the source message x . Any lower bounds on $L(N)$ and $E_{P^*, P_1^*}\{G_N^*(X | W)\}$ for this informed guesser are also lower bounds for the original, uninformed guesser because the class of guessing strategies with side information is a superset of the class of guessing strategies without it. We shall consider a specific encryption function from [21]. Let $G_N(w)$ be an arbitrary guessing strategy for this encryption function and $e(L(N), G_N(w), \Delta) \leq \exp\{-NE\}$ for some $L(N)$.

The encryption function is constructed as follows. A given source vector x is represented as a sequence (w_{b_1}, w_{b_2}) of $(b_1 + b_2)$ bits: the first $b_1 = \lceil \log |P(X, N)| \rceil$ bits of the sequence describe the index of the type P of the set $T_P(X)$ that contains x , the last $b_2 = \lceil \log |T_P(X)| \rceil$ bits identify the index of x within $T_P(X)$. The encryption function is constructed by different ways for the cases $\exp\{K\} < |T_P(X)|$ and $\exp\{K\} \geq |T_P(X)|$.

1) Consider the case $\exp\{K\} < |T_P(X)|$.

The set $T_P(X)$ is partitioned into $M = \lfloor |T_P(X)| / \exp\{K\} \rfloor$ disjoint subsets $T_P^{(1)}(X)$, $T_P^{(2)}(X)$, ..., $T_P^{(M)}(X)$, each of size $\exp\{K\}$ and if $|T_P(X)| / \exp\{K\}$ is not integer an additional subset $T_P^{(M+1)}(X)$. Then the index of the subset $T_P^{(m)}(X)$ ($1 \leq m \leq M+1$) which contains x is identified by a sequence w_{c_1} of $c_1 = \log(M+1) = \lceil \log |T_P(X)| \rceil - K$ bits, and the index of x within $T_P^{(m)}(X)$ - by a sequence of $c_2 = K$ bits. Let the encryption function f be following: $f(x, u) = (w_{b_1}, w_{c_1}, w_{c_2})$, where the sequence w_{c_2} of last c_2 bits is obtained after using simple bit-by-bit XOR operation on the c_2 bits describing the index of x within $T_P^{(m)}(X)$ and on the bits of u .

Let $W(x)$ be the set of all cryptograms w of the fixed vector x . Then the conditional probability of w given $x \in T_P(X)$ is equal

$$Pr(w | x) = Pr(u | x) = P_1^K(u) = \begin{cases} \exp\{-K\}, & w \in W(x), \\ 0, & \text{elsewhere.} \end{cases}$$

Using Bayes rule and (7) the conditional probability of $x \in T_P(X)$ given $w = f(x, u)$ (that is $x \in W^{-1}(w)$) is

$$\begin{aligned} Pr(x | x \in T_P(X), w) &= \frac{Pr(x | x \in T_P(X))P(w | x)}{\sum_{x' \in T_P(X)} Pr(x' | x' \in T_P(X))P(w | x')} = \\ &= \frac{\exp\{-NH_P(X) - ND(P \| P^*) - K\}}{\sum_{x' \in T_P(X) \cap W^{-1}(w)} \exp\{-NH_P(X) - ND(P \| P^*) - K\}} = |W^{-1}(w) \cap T_P(X)|^{-1}. \end{aligned} \quad (11)$$

Let $A_1(L(N), G_N(w), \Delta)$ be the set of those $x \in A(L(N), G_N(w), \Delta)$ which can be guessed successfully only by one vector from the guessing strategy $G_N(w)$:

$$A_1(L(N), G_N(w), \Delta) = \{x \in A(L(N), G_N(w), \Delta) : \forall y \in A(L(N), G_N(w), \Delta), G_N(x | w) \neq G_N(y | w)\}.$$

For brevity instead of $A(L(N), G_N(w), \Delta)$ we shall just write A , instead of $A_1(L(N), G_N(w), \Delta)$ - just A_1 .

We can write:

$$|A_1 \cap T_P(X)| = |T_P(X)| - |\overline{A_1} \cap T_P(X)|.$$

a) If $x \in T_P^{(m)}(X)$ ($1 \leq m \leq M$), then from (11)

$$Pr(x | x \in T_P^{(m)}(X), w) = |T_P^{(m)}(X)|^{-1} = \exp\{-K\}.$$

Therefore from (8) we obtain

$$|\bar{A}_1 \cap T_P(X)| = \frac{Pr(\bar{A}_1 \cap T_P(X))}{Pr(x | x \in T_P^{(m)}(X), w)} \leq \frac{P^{*N}(T_P(X))}{\exp\{-K\}} \leq \exp\{K - ND(P \| P^*)\}. \quad (12)$$

Hence for $P \neq P^*$ and N large enough

$$\begin{aligned} |\bar{A}_1 \cap T_P(X)| &\geq \exp\{K\} - \exp\{K - ND(P \| P^*)\} = \\ &= \exp\{K\}(1 - \exp\{-ND(P \| P^*)\}) \geq \exp\{K - 1\}. \end{aligned} \quad (13)$$

To find low estimate of maximum of $G_N(x | w)$ for $x \in \bar{A}_1 \cap T_P(X)$ we can write as for arithmetical progression

$$\sum_{x \in \bar{A}_1 \cap T_P(X)} G_N(x | w) \geq \sum_{i=1}^{|\bar{A}_1 \cap T_P(X)|} i = \frac{1 + |\bar{A}_1 \cap T_P(X)|}{2} \cdot |\bar{A}_1 \cap T_P(X)| \quad (14)$$

and

$$\sum_{x \in \bar{A}_1 \cap T_P(X)} G_N(x | w) < |\bar{A}_1 \cap T_P(X)| \cdot \max_{x \in \bar{A}_1 \cap T_P(X)} G_N(x | w). \quad (15)$$

Therefore from (13), (14) and (15) we obtain that for $P \neq P^*$

$$\max_{x \in \bar{A}_1 \cap T_P(X)} G_N(x | w) > \frac{1 + |\bar{A}_1 \cap T_P(X)|}{2} \geq \frac{1 + \exp\{K - 1\}}{2} \geq \exp\{K - 2\}. \quad (16)$$

b) For $x \in T_P^{(M+1)}(X)$ from (11) we receive that the conditional probability

$$Pr(x | x \in T_P^{(M+1)}(X), w) = |T_P^{(M+1)}(X)|^{-1} \geq \frac{1}{\exp\{K\} - 1}.$$

Then from (8) the following estimates take place:

$$\begin{aligned} |\bar{A}_1 \cap T_P(X)| &= \frac{Pr(\bar{A}_1 \cap T_P(X))}{Pr(x | x \in T_P^{(M+1)}(X), w)} \leq P^{*N}(T_P(X)) \cdot (\exp\{K\} - 1) \leq \\ &\leq \exp\{-ND(P \| P^*)\} \cdot (\exp\{K\} - 1). \end{aligned}$$

Therefore for $P \neq P^*$ and N large enough

$$\begin{aligned} |\bar{A}_1 \cap T_P(X)| &\geq \exp\{K\} - \exp\{-ND(P \| P^*)\} \cdot (\exp\{K\} - 1) \geq \\ &\geq \exp\{K\}(1 - \exp\{-ND(P \| P^*)\}) + \exp\{-K - ND(P \| P^*)\} \geq \exp\{K - 1\}. \end{aligned} \quad (17)$$

From (14), (15) and (17) we receive that (16) holds. Remark that from definition of \bar{A}_1 it follows that

$$\max_{x \in \bar{A}_1 \cap T_P(X)} G_N(x | w) = \max_{x \in \bar{A}_1 \cap T_P(X)} G_N(x | w). \quad (18)$$

Hence it must be

$$L(N) \geq \max_{P \in \alpha(E)} \max_{x \in A \cap T_P(X)} G_N(x | w) \geq \exp\{K - 2\}, \quad (19)$$

from where we can receive one of the estimates for R'' in (1).

Now for bounding of R we estimate

$$\begin{aligned} E_{P^*, P_1^*} \{G_N(X | w)\} &= \sum_{x \in \mathcal{X}^N} \sum_{u \in \mathcal{U}^K} P^{*N}(x) P_1^{*K}(u) G_N(x | w) \geq \\ &\geq \sum_{P \in \alpha(E)} \sum_{x \in T_P(X)} \sum_{u \in \mathcal{U}^K} P^{*N}(x) P_1^{*K}(u) G_N(x | w) \geq \\ &\geq \max_{P \in \alpha(E)} \sum_{x \in T_P(X)} \sum_{u \in \mathcal{U}^K} P^{*N}(x) P_1^{*K}(u) G_N(x | w). \end{aligned} \quad (20)$$

Using (5) we continue the estimates (12)

$$\begin{aligned} |\overline{A}_1 \cap T_P(X)| &\leq \exp\{K - ND(P \| P^*)\} \leq |T_P(X)| \exp\{ND(P \| P^*)\} \leq \\ &\leq \exp\{N(H_P(X) - D(P \| P^*))\}. \end{aligned} \quad (21)$$

Therefore from (5) for N large enough

$$\begin{aligned} |A_1 \cap T_P(X)| &= |T_P(X)| - |\overline{A}_1 \cap T_P(X)| \geq \\ &\geq \exp\{NH_P(X)\} \left((N+1)^{-|x|} - \exp\{-ND(P \| P^*)\} \right) = \\ &= (N+1)^{-|x|} \exp\{NH_P(X)\} \left(1 - (N+1)^{|x|} \exp\{-ND(P \| P^*)\} \right) \geq \\ &\geq \frac{1}{2} (N+1)^{-|x|} \exp\{NH_P(X)\} \geq \exp\{N(H_P(X) - \epsilon)\}. \end{aligned} \quad (22)$$

Taking into account (7), (17), (20) and (22) and definition of the set A_1 we obtain

$$E_{P^*, P_1^*} \{G_N(X | w)\} \geq \max_{P \in \alpha(E)} \exp\{-N(D(P \| P^*) + H_P(X))\} \sum_{i=1}^{|A_1 \cap T_P(X)|} i \sum_{u \in \mathcal{U}^K} P_1^{*K}(u).$$

Using the properties of arithmetic progression and (13) we can continue estimation

$$\begin{aligned}
 & E_{P^*, P_1^*} \{G_N(X | w)\} \geq \\
 & \geq \max_{P \in \alpha(E)} \exp\{-N(D(P \| P^*) + H_P(X))\} \frac{1 + |A_1 \cap T_P(X)|}{2} \cdot |A_1 \cap T_P(X)| \geq \\
 & \geq \max_{P \in \alpha(E)} \exp\{-N(D(P \| P^*) + H_P(X))\} \frac{|A_1 \cap T_P(X)|^2}{2} \geq \\
 & \geq \max_{P \in \alpha(E)} \exp\{-N(D(P \| P^*) + H_P(X))\} \exp\{K - 2\} \exp\{N(H_P(X) - \varepsilon)\} = \\
 & = \max_{P \in \alpha(E)} \exp\{-ND(P \| P^*)\} \exp\{K - 2\} \exp\{-N\varepsilon\} \geq \\
 & \geq \max_{P \in \alpha(E)} \exp\{N(-D(P \| P^*) + R' - 2\varepsilon)\}.
 \end{aligned} \tag{23}$$

Thus we receive one of the estimates for R in (2).

2) Now consider the case $\exp\{K\} \geq |T_P(X)|$. In this case it is natural to use the following encryption function: $f(x, u) = (w_{b_1}, w_{b_2})$, where the first $b_1 = \lceil \log |\mathcal{P}(\mathcal{X}, N)| \rceil$ bits (the bits of w_{b_1}) describe the index of the set $T_P(X)$ containing x , the last $b_2 = \lceil \log |T_P(X)| \rceil$ bits (the bits of w_{b_2}) are obtained after encryption by simple bit-by-bit XOR operation on the b_2 bits identifying the index of x within $T_P(X)$ and on the bits of u .

Then the conditional probability of w given $x \in T_P(X)$ is equal

$$\Pr(w | x) = \begin{cases} \exp\{-\lceil \log |T_P(X)| \rceil\}, & w \in \mathcal{W}(x), \\ 0, & \text{elsewhere.} \end{cases}$$

By analogy with (11) and using (5) we have

$$\Pr(x | x \in T_P(X), w) = |\mathcal{W}^{-1}(w) \cap T_P(X)|^{-1} = |T_P(X)|^{-1} \geq \exp\{-NH_P(X)\}.$$

Then for $P \in \alpha(E)$

$$|\bar{A} \cap T_P(X)| = \frac{\Pr(\bar{A} \cap T_P(X))}{\Pr(x | x \in T_P(X), w)} \leq \exp\{N(H_P(X) - E)\},$$

therefore for sufficiently large N and $0 < \varepsilon < E$ we can have the following lower estimate

$$\begin{aligned}
 |A \cap T_P(X)| & \geq (N+1)^{-|x|} \exp\{NH_P(X)\} - \exp\{N(H_P(X) - E)\} = \\
 & = \exp\{N(H_P(X) - \varepsilon)\} (\exp\{N\varepsilon\} \cdot (N+1)^{-|x|} - \exp\{N(\varepsilon - E)\}) \geq \\
 & \geq \exp\{N(H_P(X) - \varepsilon)\} (2 - \exp\{N(\varepsilon - E)\}) \geq \exp\{N(H_P(X) - \varepsilon)\}.
 \end{aligned} \tag{24}$$

Now we shall obtain upper estimate for $|A \cap T_P(X)|$. To each $x \in A \cap T_P(X)$ an unique guessing vector $\hat{x}_i(w) \in \mathcal{G}_N(w)$ corresponds, such that $G_N(x | w) = i$. This vector determines a conditional type $Q = Q_P(\hat{x}(w) | x)$, for which $\hat{x}_i(w) \in T_{P,Q}(\hat{X}(w) | x)$. Remark that for

the case $K \geq \log |\mathcal{T}_P(X)|$ (according to the definition of the constructed encryption function f) the wiretapper is informed only about the type $\mathcal{T}_P(X)$ of the source message \mathbf{x} . Therefore

$$Q = Q_P(\hat{x}(w) | x) = Q_P(\hat{x} | x) = Q_P \text{ and } \mathcal{T}_{P,Q}(\hat{X}(w) | x) = \mathcal{T}_{P,Q_P}(\hat{X} | x).$$

Since $\mathbf{x} \in \mathcal{A}$, then $E_{P,Q_P} d(X, \hat{X}(w)) = d(\mathbf{x}, \hat{\mathbf{x}}_i(w)) \leq \Delta$. So, $Q_P \in \mathcal{M}(P, \Delta)$. The set of all vectors $\mathbf{x} \in \mathcal{A} \cap \mathcal{T}_P(X)$ is divided into classes corresponding these conditional types Q_P and is the union of all such Q_P - shells. Let us select $Q_P = \Phi(P)$ - shell having maximal cardinality for given P and denote it by $(\mathcal{A} \cap \mathcal{T}_P(X)) (\Phi(P))$. Using (4) we have for N large enough

$$\begin{aligned} |\mathcal{A} \cap \mathcal{T}_P(X)| &\leq (N+1)^{|\mathcal{X}| |\hat{\mathcal{X}}|} |(\mathcal{A} \cap \mathcal{T}_P(X)) (\Phi(P))| \leq \\ &\leq \exp\{N\epsilon\} |(\mathcal{A} \cap \mathcal{T}_P(X)) (\Phi(P))|. \end{aligned} \quad (25)$$

Let $\mathcal{C}(P, \Phi(P), \mathcal{A})$ be the set of all guessing vectors $\hat{\mathbf{x}}_i(w) \in \mathcal{G}_N(w)$, which satisfy $G_N(\mathbf{x} | w) = i$ for some $\mathbf{x} \in \mathcal{A} \cap \mathcal{T}_P(X)$, $\mathbf{x} \in \mathcal{T}_{P, \Phi(P)}(X | \hat{\mathbf{x}}_i(w))$. In accordance with the definition of the guessing strategy $\mathcal{G}_N(w)$

$$|\mathcal{C}(P, \Phi(P), \mathcal{A})| \leq \max_{\mathbf{x} \in \mathcal{A} \cap \mathcal{T}_P(X)} G_N(\mathbf{x} | w).$$

Then

$$\begin{aligned} |(\mathcal{A} \cap \mathcal{T}_P(X)) (\Phi(P))| &\leq \sum_{\hat{\mathbf{x}} \in \mathcal{C}(P, \Phi(P), \mathcal{A})} |\mathcal{T}_{P, \Phi(P)}(X | \hat{\mathbf{x}}(w))| \leq \\ &\leq \exp\{NH_{P, \Phi(P)}(X | \hat{X})\} \cdot \max_{\mathbf{x} \in \mathcal{A} \cap \mathcal{T}_P(X)} G_N(\mathbf{x} | w). \end{aligned}$$

From the last inequality, (24) and (25) we obtain that for $P \in \alpha(E)$

$$\max_{\mathbf{x} \in \mathcal{A} \cap \mathcal{T}_P(X)} G_N(\mathbf{x} | w) \geq \exp\{N(I_{P, \Phi(P)}(X \wedge \hat{X}) - 2\epsilon)\}. \quad (26)$$

Hence the estimate of the number of guesses $L(N)$ must be

$$L(N) \geq \exp\{N(R(P^*, E, \Delta) - 2\epsilon)\}. \quad (27)$$

Using (8) we receive (21):

$$\begin{aligned} |\overline{\mathcal{A}}_1 \cap \mathcal{T}_P(X)| &= \frac{\Pr(\overline{\mathcal{A}}_1 \cap \mathcal{T}_P(X))}{\Pr(\mathbf{x} | \mathbf{x} \in \mathcal{T}_P(X), w)} \leq \frac{P^{*N}(\mathcal{T}_P(X))}{\exp\{-NH_P(X)\}} \leq \\ &\leq \exp\{N(H_P(X) - D(P \| P^*))\}, \end{aligned}$$

hence (22) too. Remark that

$$|\mathcal{A}_1 \cap \mathcal{T}_P(X)| \leq |\mathcal{T}_P(X)| \leq \exp\{NH_P(X)\}. \quad (28)$$

Using (7), (18), (20), (22), (26), (28) and definition of the set \mathcal{A}_1 we obtain

$$\begin{aligned}
 E_{P^*, P_1^*} \{G_N(X | w)\} &\geq \max_{P \in \mathcal{O}(\mathcal{E})} \sum_{x \in \mathcal{A}_1 \cap T_P(X)} P^{*N}(x) \sum_{u \in I^K} P_1^{*K}(u) G_N(x | w) \geq \\
 &\geq \max_{P \in \mathcal{O}(\mathcal{E})} \exp\{-N(D(P \| P^*) + H_P(X))\} \left(\max_{x \in \mathcal{A}_1 \cap T_P(X)} G_N(x | w) + \frac{|\mathcal{A}_1 \cap T_P(X)| - 1}{\sum_{i=1}^K i} \right) \geq \\
 &\geq \max_{P \in \mathcal{O}(\mathcal{E})} (\exp\{-N(D(P \| P^*) + H_P(X))\} \cdot \\
 &\cdot (\exp\{N(I_{P, \Phi(P)}(X \wedge \hat{X}) - 2\varepsilon)\} + (|\mathcal{A}_1 \cap T_P(X)|^2 - |\mathcal{A}_1 \cap T_P(X)|)/2)) \geq \\
 &\geq \max_{P \in \mathcal{O}(\mathcal{E})} (\exp\{-N(D(P \| P^*) + H_P(X))\} \cdot \\
 &\cdot (\exp\{N(I_{P, \Phi(P)}(X \wedge \hat{X}) - 2\varepsilon)\} + \exp\{2N(H_P(X) - \varepsilon) - 1\} - \exp\{NH_P(X) - 1\})) \geq \\
 &\geq \max_{P \in \mathcal{O}(\mathcal{E})} (\exp\{N(-D(P \| P^*) + I_{P, \Phi(P)}(X \wedge \hat{X}) - 2\varepsilon)\} \cdot \\
 &\cdot (\exp\{-NH_P(X)\} + \exp\{N(H_{P, \Phi(P)}(X | \hat{X}) - \varepsilon)\} - \exp\{-N(I_{P, \Phi(P)}(X \wedge \hat{X}) - \varepsilon)\})) \geq \\
 &\geq \max_{P \in \mathcal{O}(\mathcal{E})} \exp\{N(-D(P \| P^*) + I_{P, \Phi(P)}(X \wedge \hat{X}) - 2\varepsilon)\}.
 \end{aligned} \tag{29}$$

Therefore uniting the both cases, from (19) and (27) we receive

$$L(N) \geq \exp\{\min(K - 2, N(R(P^*, E, \Delta) - 2\varepsilon))\},$$

from (23) and (29) we obtain

$$E_{P^*, P_1^*} \{G_N(X | w)\} \geq \max_{P \in \mathcal{O}(\mathcal{E})} \exp\{N(-D(P \| P^*) + \min(R' - 2\varepsilon, I_{P, \Phi(P)}(X \wedge \hat{X}) - 2\varepsilon))\}.$$

Hence for N large enough

$$R'' \geq \frac{1}{N} \log L(N) - \varepsilon \geq \min(R' - \varepsilon, R(P^*, E, \Delta) - 3\varepsilon),$$

$$R \geq \frac{1}{N} \log E_{P^*, P_1^*} \{G_N(X | w)\} - \varepsilon \geq$$

$$\geq \max_{P \in \mathcal{O}(\mathcal{E})} (-D(P \| P^*) + \min(R' - 2\varepsilon, R(P, \Delta) - 2\varepsilon)) - \varepsilon.$$

Taking into account arbitrariness of ε , continuity by E of all functions in above expressions, we obtain (1) and (2), therefore the proof of the inclusion (10) is completed.

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Կոստիս ըստ շեղման և հուսալիության չափանիշների Շենոնի ծածկագրման համակարգի համար

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Անիմիտ

Լուծվում է դիսկրետ առանց հիշողության աղբյուրի հաղորդագրությունների արված շեղման և հուսալիության մակարդակի սահմաններում կոստիս խնդիրը Շենոնի ծածկագրման համակարգի համար: Գաղափարները չափվում է աղբյուրի ծածկագրի հիման վրա կատարվող հակառակորդի հաղորդագրությունների վերականգնման գույակումների միջին քվո: Ի հավելում Սերիալի և Արիկանի կողմից դիտարկված խնդիր, աշխատանքում դիտարկվում են շեղման և հուսալիության լրացուցիչ պայմանները:

Բանալտ արագության համար որոշված է շեղման և հուսալիության պահանջվող մակարդակներն ապահովող կոստիսն փորձերի փոքրագույն (ըստ բոլոր գույակման ցուցանիշների) և մեծագույն (ըստ բոլոր ծածկագրման ֆունկցիաների) քվի համապատասխան կոստիսների ապահովի արժեքը: